The impact of selected factors and access to mineral resources on the development of wind energy in Poland

Introduction

Decarbonization resulting from the implementation of the European Green Deal requires the introduction of extensive changes in the Polish energy mix. Currently, the basis of the energy generation structure in Poland is hard coal and lignite, which in 2022 was used to generate 71% of the consumed electricity (BP 2023). In 2022, the potential of generating capacity of this type was approximately 7,000 MW, and the annual production of wind energy oscillated around 8 GWh. This is just over 10% of the annual energy demand in Poland. The Polish Energy Policy assumes that by 2040, approximately 50% of electricity in Poland will be obtained from renewable energy sources (PEP 2040). These will be primarily wind and solar energy. In 2023, the installed capacity of onshore wind farms
in Poland was 8.6 GW. The policy assumes the further development of wind energy from not only onshore but also offshore farms. By 2030, their installed capacity is to reach about 6 GW, and by 2040 it is to reach 11 GW. Therefore, by 2040, the capacity of Polish wind farms will most likely double. Poland’s energy policy until 2040 assumes also that the role of hard coal in the economy will gradually decrease. Reducing the emission intensity of the energy sector will be possible through, among other means, the expansion of wind energy and the implementation of offshore wind energy. Renewable energy sources will take over an increasingly important role in the national energy mix. By 2030, it is to be at least 32% of the gross final energy consumption. This development will mainly be based on photovoltaics and wind farms. Onshore wind farms will continue to be developed. Additionally, PEP2040 assumes the development of offshore farms. This is mainly due to the productivity of these installations and the fact that they eliminate problems with social acceptance, which is often an obstacle in the case of onshore farms. Electricity production from offshore farms will have the largest share in total production from renewable energy sources. This is also consistent with the goals of the European Green Deal, which makes the development of offshore farms a strategic project of Poland’s Energy Policy (PEP2040). Therefore, in order to replace coal, it will be necessary to significantly expand the potential of wind energy in Poland. Such an intensive evolution of the energy system will depend on many factors that will have positive or negative impacts on the pace of wind energy development. These include, among other factors, climatic and economic conditions, social moods, legal actors, and above all, the availability of critical raw materials on global markets that are necessary for the construction of wind farms. The Act of May 20, 2016 on investments in wind farms, and, above all, the 10 H rule contained therein, which determines the minimum distance of wind turbines from residential buildings, has resulted in the inhibition of the development of wind energy in Poland. In 2023, an amendment to the Wind Farm Act was introduced, where the distance guidelines were relaxed. This is a step towards the implementation of the PEP2040 assumptions, which may accelerate the development of onshore energy in Poland. The presented research focused only on quantitative factors. The following hypothesis was put forward: factors shaping the pace of the potential development of wind energy generation in Poland would allow the fulfillment of the assumptions of the Polish energy policy. Therefore, the presented research analyzes a set of factors that may potentially affect the development of wind energy in Poland. First, the focus was on critical raw materials, but factors that were included in the economic, ecological, and technological categories were also taken into account. The authors examined the impact of the factors adopted for the analysis on the volume of wind energy production in Poland. The statistical significance of the factors was verified using the ARMAX model, which was also used to build a forecast of the energy production volume in the selected time horizon.
1. Literature review

Wind energy is an inexhaustible source of energy, which in the light of rising electricity prices is an increasingly attractive solution (EON 2023). Energy production in wind farms is also free from pollutant emissions and greenhouse gases such as CO\(_2\) (Yousefi et al. 2019; Forbes and Zampelii 2019). CO\(_2\) emissions are one of the basic problems that significantly accelerated the decisions of the European Union and its member states concerning the need to introduce changes in energy mixes. In the literature, many publications on wind energy forecasting can be found (Gil et al. 2010). Usually, these are short-term analyses. Very few researchers have attempted to introduce Gray’s forecasting method, which was designed for system analysis characterized by insufficient information. An empirical study based on real data obtained from a wind farm located in Penghu, Taiwan, demonstrated the effectiveness of the GM-based forecasting mechanism. It has also been applied to wind forecasting problems at different time intervals (Huang et al. 2011). Short-term wind speed forecasting using spectral analysis was also used. Experimental results show that trimming daily, weekly, monthly, and yearly patterns in the measurements significantly increases the accuracy of the estimates. The proposed framework is based on data detrending, covariance factorization using a recently developed subspace method, and prediction concepts using a single- and/or multi-step-ahead Kalman filter (Akçay and Filik 2017). The authors demonstrated that SVR is an effective method for predicting wind energy production based solely on wind measurements from windmills, especially without further meteorological data and weather forecasts (Kramer and Gieske 2011). Machine learning was used for the short-term forecasting of wind energy generation. The results showed that the algorithm proposed by the authors has a 10% lower RMSE value compared to other models (Shabbier et al. 2019). Long-term forecasts were also made using linear and nonlinear models. Various regression models were developed using, inter alia, electricity consumption and gross domestic product (GDP). Statistical tests were used to check the validity of the proposed models. It was indicated that the developed models are consistent with official forecasts (Bianco et al. 2009; Ha et al. 2019). The trends in the share of renewable energy in final energy consumption were analyzed using data for the EU. Empirical estimates of the share of renewable energy made using ARIMA models showed an increasing trend (Mehedintu et al. 2018). The authors used economic and sociological factors as input. In forecasting the volume of wind energy production, neural networks were also used to forecast the wind speed on the horizon of six months (Azad et al. 2014). However, the problem with forecasts based on meteorological data is that they can only reach the forecast horizon of the meteorological model at most. The literature also presents forecasts with a one-year time horizon using Bayesian modelling and Markov chain Monte Carlo (MCMC). The data presented in the literature show that forecasting the volume of wind energy production is complicated due to the nature of the phenomenon. A satisfactory forecast is obtained by dividing the problem into several periods with different beta distribution parameters. As a result of this, accurate forecast can be obtained (Mesa-Jimenez et al. 2023).
REE demand scenarios resulting from the demand for wind energy have been presented in the literature. It was examined whether the supply of REE on a global scale is able to meet the needs of the developing offshore wind energy (Kalvig and Machacek 2018; Imholte at al. 2018). The demand and prices of materials and metals were also forecasted to avoid future problems with the availability of these materials. In this study, lithium and cobalt, as well as neodymium, praseodymium and dysprosium were considered to be critical materials (Verma et al. 2022).

The authors of the presented research, in turn, wanted to create a long-term forecast of the volume of the production of wind energy in the horizon of two years, which would take into account the demand for critical metals necessary during the construction of wind farms. Therefore, information on the amount of demand for copper, nickel, boron, manganese and REE (rare earth metals) was introduced into the model.

2. Critical metals characteristics

The critical metals necessary in the wind turbine manufacturing process are briefly presented below.

**Rare earth metals**

Rare earth metals are yttrium, scandium, and a group of seventeen elements from the lanthanide group (Tao et al. 2022, Rybak and Rybak 2021). Access to rare earth elements will determine the success of the energy transformation of countries around the world in the future. Because REE will have a key impact on the development of the energy sector, as well as on the development of emission-free technologies, they are called vitamins of the new industry (Balaram 2019). This also applies to wind turbines (Ballinger et al. 2020). REE resources are located mainly in China, as well as the USA, Brazil, India and Australia (Jaroni et al. 2019). Since there are no REE deposits within the geographical area of the EU that would enable meeting the demand of member states, they are considered critical raw materials (COM 2023). REEs such as neodymium, dysprosium, praseodymium, and terbium are essential to the production of wind turbines. Poland does not have its own REE deposits of which exploitation could be profitable. However, an alternative native source of elements in this case may be fly ash generated in the coal combustion process (Blissett and Rowson 2012).

**Copper**

Copper in wind turbines is used in the generator located on the nacelle, in cables and wires, transformers, grounding, inverters and lightning protection and control systems (Carrara et al. 2020; Farina and Anctil 2022). Poland has its own copper deposits that will successfully cover the needs of the wind farm development process. Polish copper deposits are exploited by KGHM Polska Miedź. They constitute one of the largest deposits in the world.
and are extracted in three mines: Rudna, Lubin and Polkowice-Sieroszowice. KGHM’s exploitable resources contain approximately 40 milligrams of copper (Geoportal 2021). In 2021, KGHM’s production of payable copper amounted to 753.7 thousand Mg (KGHM 2021), which means a 7.5% increase compared to 2019. The growing demand for this raw material, which has been taking place since the beginning of the twentieth century, is currently due to the development of technological innovations, including those related to the development of renewable energy sources. It is estimated that the demand for copper and its prices will increase in the coming years as a result of the war in Ukraine as well as the demand caused by the green transformation and the need to ensure energy security in, among other places, the European Union and the United States (GS 2022).

**Nickel**

Nickel is used in many elements of wind turbines, mainly bearings, gears, shafts and hydraulic elements of nacelles, as well as in many other structural elements of wind turbines (Lee 2008; Carrara et al. 2020). In Poland, nickel is recovered in the processing of copper-silver ores from deposits of the Fore-Sudetic Monocline.

**Boron**

Boron is used to boronize the steel elements of wind turbines, such as gears. This makes them resistant to heat, abrasion and corrosion (Greco et al. 2011). Boron is also used in the permanent magnets of the turbine generator (Carrara et al. 2020). Boron compounds used in Poland are mainly imported from Italy and Germany (Witkowska-Kita et al. 2016).

**Manganese**

Small amounts of manganese are used as part of the steel components of turbines (USAID 2021). In Poland, manganese demand is mainly covered by imports from Brazil and Ukraine (Witkowska-Kita et al. 2016).

The explanatory variables selected by the authors were introduced into the ARMAX model. In addition, during the research, the authors took into account ecological, energy and economic factors, which were also introduced into the ARMAX model. The model and the WEKR 1.0 program are characterized in the Methods and Results chapters.

In Germany, copper and dysprosium have been identified as the most critical materials because they could cause supply bottlenecks while being fundamental to the functionality of wind turbines. It was determined that the demand for copper could require 0.2% of the currently known resources, and the demand for dysprosium could reach up to 0.6% of reserve levels. Both metals clearly exceed allocations for renewable energy technologies in Germany and will face strong competition from other sectors for raw materials (Shammugan et al. 2019).

Ren, Tan et al. noted that the annual demand for base metals for Chinese wind energy in 2050 will be as much as twelve times higher than in 2018, and the cumulative demand will be as much as twenty-three times higher. The cumulative copper and nickel demand of the
wind power sector is 9–11.9 Mt and 2.1–2.8 Mt, respectively, accounting for 35–45.9% of the copper resources and 74–101% of the nickel resources in China. In terms of the demand for rare earths, the annual demand is projected to increase by more than eighteen times in 2050 compared to 2020, with cumulative demand for neodymium and dysprosium accounting for 1.6–3.3% and 1.4–2.8% of their reserves, respectively (Ren et al. 2021). Figure 1 presents global resources of selected REEs and the demand for them in the wind energy sector over the next thirty years. It is clearly visible that, especially in the case of Tb, where the demand related to wind energy will consume approximately 78% of resources by 2050, and Dy, where it will be 34%, the lack of access to new sources of REE will be a problem in the near future. It should also be noted that the demand for REE will increase in other sectors of the economy and not only in wind energy, which may lead to a complete lack of access to REE.

Copper consumption per MW can be found in many publications. For example, in (Ardente et al. 2008; Crawford 2009) it ranges from 0.36 to 1 Mg/MW. In turn, it is approximately 0.7 Mg/MW. With the average assumption based on (Alves Dias et al. 2020) taking into account only Polish Cu resources, the annual demand for copper from wind farms was approximately 0.05% of production in 2022.

The literature also includes examples of determining the share of demand for metals in connection with the development of wind energy in 2020 in annual production (Li et al. 2022). For Cu, it was 0.1%, Ni 0.2% and Mn 0.0%. The authors performed a similar analysis for Poland, the results of which are presented in Figure 2. The results of research conducted for Poland are comparable. The only discrepancy occurs in the case of Nickel, because for
Poland, only domestic production of the metal recovered during copper mining was taken into account.

3. Methods

The methodology according to which the research was performed is presented in Figure 3.

Step 1
The initial analysis took twenty-four factors into account selected on the basis of a literature review and classified into five categories:
Technological factors:
- annual installed capacity of wind energy, (MW);
- number of patents.

Energy factors:
- wind energy production, (GWh) – dependent variable;
- primary energy consumption, (GJ/capita);
- gross available energy, (toe);
- total energy supply, (toe);
- final energy consumption, (toe);
- renewables consumption, (EJ).

Ecological factors:
- CO₂ emission, (mil Mg).

Economic factors:
- real GDP per capita, (EUR/capita);
- LCOE for onshore technology, (USD/kWh).

Raw material factors:
- Nd consumption, (kg/MW);
- Pr consumption, (kg/MW);
- Tb consumption, (kg/MW);
- Dy consumption, (kg/MW);
- annual REE production, (1000 Mg);
- Ni consumption, (kg/MW);
- Br consumption, (kg/MW);
- Cu consumption, (kg/MW);
- Mn consumption, (kg/MW);
- annual domestic production of Cu (1000 Mg);
- annual domestic production of Ni, (Mg);
- annual world production of Mn, (1000 Mg);
- annual world production of Br, (1000 Mg).

Step 2
In the second step, statistical data necessary to conduct the presented research was obtained. Table 1 presents sources of data acquisition.

Step 3
The consumption of individual raw materials was determined using the WEKR 1.0 program based on data from the report (Carrara et al. 2020). Figure 4 presents an algorithm, on the basis of which the programme determines the selected explanatory variables based on the example of nickel. The programme determines the number of wind energy MW that have been created in Poland since 1998. In the next step, it calculates how many kilograms of each metal were needed in a given year to generate the wind power generation capacity.
Table 1. Data acquisition sources

<table>
<thead>
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<th>Factor</th>
<th>Source</th>
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<td>Eurostat</td>
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<tr>
<td>Number of patents</td>
<td>IRENA</td>
</tr>
<tr>
<td>Energy factors</td>
<td>Eurostat</td>
</tr>
<tr>
<td>Ecological factors</td>
<td>Eurostat</td>
</tr>
<tr>
<td>Real GDP per capita, EUR/capita</td>
<td>Eurostat</td>
</tr>
<tr>
<td>LCOE for onshore technology, USD/kWh</td>
<td>IRENA</td>
</tr>
<tr>
<td>Cu, Ni, Mn, Br consumption</td>
<td>MEERI PAS, KGHM</td>
</tr>
<tr>
<td>Nd, Pr, Tb Dy consumption</td>
<td>MEERI PAS</td>
</tr>
</tbody>
</table>

Fig. 4. The algorithm of the WEKR 1.0 program
Source: own elaboration

Rys. 4. Algorytm program WEKR 1.0
The input data to the algorithm are:
- The amount of energy production,
- The number of MW per year,
- The demand for REE, Cu, Ni, Br and Mn metals.

Input data is placed in arrays. The volume of MW built each year is determined. Additionally, the demand for a given metal is determined for each year. After determining this value for all observations of the time series, the program ends. The output data, i.e. the demand for nickel in accordance with the designated production capacity expected for a given year, is printed in the console and placed in a table, which then constitutes input data for the ARMAX model.

Steps 4 and 5

The Autoregressive Moving Average with Exogenous Input model (ARMAX) was used to forecast the volume of wind energy production in Poland. The model includes a moving average component and an autoregressive component. The ARMAX model enables the examination of the relationship between the explained variable and the explanatory variables. Optimization of the model parameters is performed by limiting the mean square error.

ARMAX is a discrete model of the input-output type used in the case of stochastic processes (Darbellay and Slama 2000; Hickey et al. 2012). The ARMAX model is described by the following equation:

\[
y(t) = \sum_{i=1}^{n_x} a_i y(t-i) + \sum_{j=1}^{n_u} b_j u(t-j) + \sum_{k=1}^{n_e} c_k e(t-k) + e(t)
\]  

- \(y(t)\) – output signal sequence,
- \(t\) – time,
- \(a_i, b_j, c_k\) – prediction coefficients,
- \(e(t)\) – white noise,
- \(u(t)\) – input signal string,
- \(n\) – predictor order,
- \(i, j, k\) – lag.

As a result of the use of the ARMAX model, it is also possible to determine the nature of the explanatory variable, i.e. to indicate whether it is:
- nominant, the character of which depends on its nominal value;
- stimulant, the increase of which has a positive effect on the dependent variable;
- destimulant, the decrease of which has a positive effect on the dependent variable (Shader et al. 2003).

In the case of a model with numerous explanatory variables, it should be verified that the multicollinearity phenomenon does not occur (Niu and Lee 2022). This is present when
the variables are strongly correlated with each other. This may have a negative impact on
the ARMAX model results. To verify whether there is multicollinearity in a selected set
of variables, the VIF (variance inflation factor) can be used, which indicates how much
the variance of the estimated model coefficient increased due to collinearity with the other
explanatory variables. If the collinearity phenomenon occurs in the model, the variables do
not carry information about the influence of the independent variable on the dependent var-
iable. The VIF coefficient is determined according to the following formula (Belsley 1980;
Miles 2014):

\[
VIF_j = \frac{1}{1 - R_j^2}
\]  

\( R_j^2 \) – determination coefficient.

Therefore, in the case of the occurrence of the multicollinearity phenomenon, it should
be eliminated. This can be done in several ways. One of these is the transformation of the
independent variables. This transformation consists of expressing the explanatory variables
between which the collinearity occurred as a composite of these variables and expressing
their mutual relation to each other. Another way is to modify the set of independent variables
by eliminating those that cause the phenomenon of collinearity or by introducing additional
explanatory variables (Wetherill 1986). However, this requires changing the research as-
sumptions of the conducted and often verifying the hypotheses.

The stationarity of the time series was verified using the Dickey-Fuller (DF) test, which
was first presented in 1979. The test checks the time series for the presence of a unit root.
In this case, two hypotheses are put forward:

- H0: there is a unit root in the time series, \( \delta = 0 \);
- H1: the time series is stationary, \( \delta < 0 \).

The time series to be introduced into the ARMAX model should be stationary or should
be reduced to a stationary form. Otherwise, the model may not be able to faithfully represent
the fundamental data patterns and the results will be erroneous.

A stationary time series is characterized by a constant mean and autocorrelation over
time. One way to bring a time series to a stationary form is to differentiate it.

The Dickey-Fuller test statistic for the existence of a unit root takes the form:

\[
DF = \frac{\delta}{S(\delta)}
\]  

\( \delta \) – deterministic trend coefficient.
ARMAX validation also consisted of the selection of models characterized by the lowest value of information criteria such as Schwarz (BIC), Hannan-Quinn (HQ) and Akaike (Piłatowska 2010):

\[
BIC = -2\ln L(\hat{\theta}) + K \ln(n)
\]  
(4)

\[
HQ = -2\ln L(\hat{\theta}) + 2K \ln(n)
\]  
(5)

\[
AIC = -2\ln L(\hat{\theta}) + 2K
\]  
(6)

\( n \) – number of observations;
\( K \) – number of model parameters;
\( L(\hat{\theta}) \) – the credibility function of the model corrected by the penalty function – the function of \( K \).

Depending on the “penalty”, the following models are considered appropriate:

- With a small penalty – a model with more parameters should be selected.
- With a large penalty – a sparingly parameterized model should be selected.

The next stage of validation was to determine model errors such as MAPe, RMSe, MAE, and Theil’s Inequality Coefficient (Willmott and Matsuura 2005). The determination of several meters enables the selection of the appropriate model.

Mean absolute percentage error (MAPE) (Bliemel 1973; Farnum and Stanton 1989):

\[
MAPE = \frac{1}{n} \sum_{i=1}^{n} \left| \frac{e_i}{y_i} \right|
\]  
(7)

\( y_i \) – value of the dependent variable in period \( t \),
\( e_i \) – prediction error.

Root mean square error (RMSE) (Safi and Zeroual 2002; Almorox et al. 2005):

\[
RMSE = \sqrt{\frac{1}{n} \sum_{i=1}^{n} e_i^2}
\]  
(8)

Mean absolute error (MAE) (Chai and Draxler 2014):
Theil’s inequality coefficient (Kufel 2004):

\[ U = \frac{RMSE}{\sqrt{\frac{1}{n} \sum_{i=1}^{n} \hat{y}_i^2} + \sqrt{\frac{1}{n} \sum_{i=1}^{n} (\hat{y}_i - \hat{y}_t)^2}} \]  

\( \hat{y}_t \) — forecast of the dependent variable in period \( t \).

The residuals of the selected model were also analyzed. It was verified whether the residuals are normally distributed and whether they show autocorrelation.

The Ljung box test (Burns 2002) was used to verify the autocorrelation of residuals:

\[ Q = n(n+2) \sum_{k=1}^{h} \frac{\hat{\rho}_k^2}{n-k} \]  

\( \hat{\rho}_k^2 \) — sample autocorrelation at lag \( k \),  
\( h \) — number of lags tested.

The normality of the distribution of model residuals was confirmed using the Door-  

\[ DH = \left[ Z(\sqrt{b_1}) \right]^2 + [z_2]^2 \]  

\( Z(\sqrt{b_1}) \) — transformed sample skewness,  
\( z_2 \) — Wilson-Hilferty transformation.

4. Results

During the determination of the ARMAX model, 39 function evaluations and 104 gradient evaluations were made. Standard errors of the parameters were based on a Hessian matrix.

The multicollinearity analysis of the variables showed that they are characterized by strong collinearity. This mainly concerned the demand for rare earth elements as well as nickel, boron, manganese and copper. When independent variables are correlated, it means
that changes in one variable are related to changes in another variable. The amount of demand for metals is directly dependent on the amount of MW of wind energy generated each year, so they are correlated. In this case, the variables must be discarded or transformed. In the analyzed case, eliminating the explanatory variables was not an option so they had to be transformed (Kumari 2008). One way to do this is to sum the values of the variables and divide them by a common factor, which was necessary in the case of REE. The demand for individual REEs was summed and divided by the production per REE in a given year. The remaining variables were transformed separately. The demand for Cu, Mn, Ni and Br was divided by the production volume of these metals in a given year (Paul 2006), which was a sufficient measure. The multicollinearity phenomenon has been removed.

The multicollinearity testing was performed again. The obtained results are presented in Table 2.

According to the Belsley-Kuh-Welsch method, cond ≥ 30 means a strong almost linear correlation, while cond between 10 and 30 means a low correlation. Estimated parameters for which the variance is related to high status indicator values can always be considered to be problematic. In the tested case:

- The number of status indicators for cond ≥ 30 is 0.
- The number of status indicators for cond ≥ 10 is 1.

Values greater than 15 indicate a possible multicollinearity problem, while values greater than 30 should be expected to cause serious multicollinearity problems (IBM 2021). In the analyzed case, only one conditional index reaches a value greater than 10, which is equal to 13.79. This value is still lower than 15; therefore, the introduced modifications of the explanatory variables brought the intended result. Only in one case is there a low linear correlation, which did not affect the quality of the model.

The time series was analyzed for stationarity. For this purpose, the extended Dickey-Fuller test was performed. The null hypothesis was put forward that a zero root appears in the time series. The test confirmed the validity of the hypothesis. Therefore, it was necessary to differentiate the time series twice.

- First differentiation:
  model: \((1 - L)y = b0 + (a - 1) \times y(-1) + ... + e\)
  the estimated value \((a - 1)\) is: \(-0.588709\)
  test statistic: \(\text{tau}_c (1) = -1.8106\)
  asymptotic \(p\)-value = 0.3758
  autocorrelation of first-order residuals: 0.021

- Double differentiation:
  model: \((1 - L)y = b0 + (a - 1) \times y(-1) + ... + e\)
  the estimated value \((a - 1)\) is: \(-1.85095\)
  test statistic: \(\text{tau}_c (1) = -2.98582\)
  asymptotic \(p\)-value = 0.03625
  autocorrelation of first-order residuals: \(-0.030\)

The probability of \(p < 0.05\) means that the time series was brought in a stationary form.
Table 2. Diagnostics of collinearity acc. Belsley-Kuh-Welsch (BKW 1980)

Tabela 2. Diagnostyka kolinearności według Belsleya-Kuh-Welscha

<table>
<thead>
<tr>
<th>Lambda</th>
<th>Cond</th>
<th>Phi 1</th>
<th>Theta 1</th>
<th>Theta 2</th>
<th>REE ratio</th>
<th>REE production</th>
<th>Patents</th>
<th>Capacity</th>
<th>LCOE</th>
<th>GDP</th>
<th>E consumption</th>
<th>CO₂ emission</th>
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<td>0.013</td>
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<td>0.002</td>
<td>0.001</td>
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<td>0.003</td>
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<td>0.011</td>
<td>0.017</td>
<td>0.003</td>
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<td>0.004</td>
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<td>0.610</td>
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<td>0.003</td>
<td>0.001</td>
<td>0.002</td>
<td>0.001</td>
<td>0.001</td>
<td>0.031</td>
<td>0.121</td>
<td>0.001</td>
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<tr>
<td>0.072</td>
<td>7.826</td>
<td>0.000</td>
<td>0.000</td>
<td>0.340</td>
<td>0.140</td>
<td>0.072</td>
<td>0.019</td>
<td>0.142</td>
<td>0.129</td>
<td>0.138</td>
<td>0.137</td>
<td>0.144</td>
<td>0.134</td>
<td>0.025</td>
<td>0.542</td>
<td>0.048</td>
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<tr>
<td>0.023</td>
<td>13.794</td>
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<td>0.001</td>
<td>0.639</td>
<td>0.682</td>
<td>0.301</td>
<td>0.112</td>
<td>0.467</td>
<td>0.621</td>
<td>0.681</td>
<td>0.675</td>
<td>0.711</td>
<td>0.638</td>
<td>0.046</td>
<td>0.005</td>
<td>0.136</td>
</tr>
</tbody>
</table>

phi – AR input argument; theta – MA input argument; cond – Belsley-Kuh-Welsch condition index; Lambda – eigenvalues of the inverse covariance matrix (the smallest is 0.0230326); REE, Mn, Ni, Br, Cu ratio – the ratio of demand to metal production.
All explanatory variables from the technological, economic, ecological, energy and raw material categories were introduced to the ARMAX model. Taking into account the optimal values of ex post error and information criteria, the ARAMX (1, 2, 2) model was finally selected. The values of the calculated indicators are presented in Table 3.

Table 3. The values of the calculated indicators

<table>
<thead>
<tr>
<th>Index</th>
<th>Index value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bayesian information criterion</td>
<td>318</td>
</tr>
<tr>
<td>Hannan Quinn information criterion</td>
<td>305</td>
</tr>
<tr>
<td>Akaike information criterion</td>
<td>303</td>
</tr>
<tr>
<td>MAE</td>
<td>207</td>
</tr>
<tr>
<td>RMS</td>
<td>252</td>
</tr>
<tr>
<td>MAPE, %</td>
<td>3</td>
</tr>
<tr>
<td>R2</td>
<td>0.99</td>
</tr>
<tr>
<td>Theil’s coefficient, %</td>
<td>0.10</td>
</tr>
</tbody>
</table>

Source: own elaboration.

Explanatory variables were verified for statistical significance. Those that were identified as statistically insignificant were removed from the ARMAX model. Ultimately, twelve variables were used in the research. Their statistical significance was verified based on the value of the Student’s t-statistic and the p value level. If $p \leq \alpha$, the null hypothesis of the variable significance should be rejected and the alternative hypothesis of the variable’s significance should be accepted. If $p > \alpha$, there is no reason to reject the null hypothesis of non-significance. In all cases, the p value was lower than the assumed significance level $\alpha = 0.01 (***)$. The factors used in the construction of the ARMAX model are presented in Table 4.

In the last column of Table 4, the nature of the variable is specified. Four factors turned out to be destimulants, which means that their increase negatively affects the volume of wind energy production. These are:

- Increase in demand for REE, which means that REE deposits may be depleted faster than expected.
- A decline in REE production would inhibit the development of wind technology.
- CO$_2$ emissions, the increase of which would mean a slowdown in the development of wind energy.
- Mn/MW, the global consumption of which, unlike Cu, Br, and Ni, has been growing intensively over the last twenty years.
Using the ARMAX model, a wind energy forecast was then made for production until 2025, which is presented in Figure 5.

The forecast indicates that by 2025 a dynamic increase in wind energy production can be expected. Considering the influence of the explanatory variables on the volume of wind energy production, assuming that their development tendency does not change, this increase may even be 50% per year. The increase is a continuation of the trend that has been observed in the volume of wind energy production in recent years. Over a ten-year period, this production increased ten times. Because this is close to exponential growth, the model maintained this trend in the forecasts until 2025. Additionally, the growth was stimulated by explanatory variables, particularly GDP, metal consumption and LCOE. In the years 1990–2015, Poland was in second place among the EU countries in the ranking of the fastest developing countries. Over the last thirty years, Poland has achieved a GDP growth of 117% per capita (RNP 2023). PEP2040 assumes that offshore wind farms will be built in the coming years.

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Parameter value</th>
<th>Standard error</th>
<th>Probability</th>
<th>Character of the variable</th>
</tr>
</thead>
<tbody>
<tr>
<td>phi_1</td>
<td>−0.96</td>
<td>0.05</td>
<td>6.56e-089 ***</td>
<td>−</td>
</tr>
<tr>
<td>theta_1</td>
<td>−1.94</td>
<td>0.24</td>
<td>2.77e-016 ***</td>
<td>−</td>
</tr>
<tr>
<td>theta_2</td>
<td>0.99</td>
<td>0.22</td>
<td>7.47e-06 ***</td>
<td>−</td>
</tr>
<tr>
<td>REE/MW</td>
<td>−0.11</td>
<td>0.00</td>
<td>0.00 ***</td>
<td>destimulant</td>
</tr>
<tr>
<td>REE production, 1000 Mg</td>
<td>−18.30</td>
<td>0.03</td>
<td>0.00 ***</td>
<td>destimulant</td>
</tr>
<tr>
<td>Patents</td>
<td>20.97</td>
<td>0.03</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>Wind energy capacity (MW)</td>
<td>0.75</td>
<td>0.01</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>LCOE (USD/kWh)</td>
<td>76,418.7</td>
<td>35.82</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>Real GDP (EUR/capita)</td>
<td>1.31</td>
<td>0.00</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>Primary energy consumption (GJ/capita)</td>
<td>111.21</td>
<td>0.05</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>CO₂ emission (mil Mg)</td>
<td>−0.87</td>
<td>0.03</td>
<td>0.00 ***</td>
<td>destimulant</td>
</tr>
<tr>
<td>Ni/MW</td>
<td>210.18</td>
<td>0.09</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>br/MW</td>
<td>3,866,250,000</td>
<td>3,509,890.00</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>Cu/MW</td>
<td>14,915,700</td>
<td>8,148.58</td>
<td>0.00 ***</td>
<td>stimulant</td>
</tr>
<tr>
<td>Mn/MW</td>
<td>−1.33e+08</td>
<td>105,164</td>
<td>0.00 ***</td>
<td>destimulant</td>
</tr>
</tbody>
</table>

Source: own elaboration.
which provides an opportunity to maintain the trend of the increasing development of wind energy generation capacity.

It should be noted that all explanatory variables from the raw materials category were considered to be factors that have a significant impact on the volume of the production of wind energy. Copper and nickel can be obtained from deposits located in the geographical area of Poland. However, boron, manganese and, above all, REE are obtained by import. The most serious bottleneck in the green transformation of the energy systems of EU member states may be REEs. Most deposits of raw materials classified as REE are located in China, which accounts for approximately 60% of the total world production (Daigle and DeCarlo 2021), as well as the United States, Brazil, Australia, CIS countries and India (Jaroni et al. 2019). REE resources located in the EU do not even reach 0.5% of global resources; therefore, the European Union is forced to import REE. This constitutes a huge threat to the energy security of the EU, which in the future wants to base its energy mix on renewable energy sources. In the European Union, several scenarios for solving this problem are being considered. Actions have been taken to diversify the sources of REE (Kalvig and Machacek 2018). The USA will become a new source of REE in the future. They decided to increase REE extraction in order to strengthen their position on world markets, which they ceded to
China in the nineteen-eighties. Another solution may be obtaining REE from recycling and coal combustion waste (Rybak and Rybak 2021). As noted, fly ash produced from Polish coal is a rich source of REE. Additionally, this method of obtaining rare earth elements eliminates the need to exploit these raw materials and, at the same time, the associated negative impact of exploitation on the natural environment (Franus et al. 2015).

It should be remembered that factors related to social acceptance will have a huge impact on the development of wind energy in Poland. In general, wind energy is perceived as an excellent way to diversify Poland’s energy mix and ensure energy security. Wind energy is perceived as accessible, free and zero-emission (Mroczek 2011). Problems with social acceptance appear at the local level when an investment is to be implemented in a specific selected location. As a result of the incorrect actions of companies erecting wind farms, turbines are associated with noise and limited visual values of the landscape (Łucki and Misiak 2011). It is necessary to undertake information and educational activities that will overcome social resistance to wind farms. Appropriate spatial planning is also important, which will enable the establishment of farms in places where they will not be in any way burdensome to the surrounding residential buildings. Another important factor is the state’s policy towards renewable energy sources, in particular wind energy. The Wind Act of 2016 effectively limited investments related to wind energy. In 2023, the regulations regarding the distance of wind turbines from buildings were relaxed, but the very fact of legislative instability may affect the future investment decisions of energy companies and private investors.

Conclusions

Wind energy in Poland currently covers only 10% of the electricity demand. To meet the assumptions of the Polish Energy Policy until 2040, it is necessary to take steps that will accelerate the dynamics of wind energy development in our country. To be effective, these actions must identify factors that will have a key impact on the building potential of the wind energy generation capacity. For this purpose, the authors built a set of variables that they considered potentially relevant in the context of this development. To confirm these assumptions, they used the ARMAX (1, 2, 2) model. As a result of this, it was possible to identify statistically significant explanatory variables, i.e. factors affecting the volume of wind energy production in Poland. Using the applied method, it was also possible to determine the nature of the impact of the individual variables on the volume of wind energy production. The forecast showed that if the development trends of the factors that affect wind energy do not change, it will be possible to meet the assumptions of PEP2040 regarding the dynamic development of wind farms in Poland and double the generation capacity by 2030. Analysis using the ARMAX model showed that access to raw materials such as REE, Cu, Ni, Br, and Mn will have a very significant impact on the development of wind energy in Poland. Each factor of the raw material category of raw materials that was introduced into the model was considered statistically significant at the significance level of $\alpha = 0.01$, i.e., at the lowest
acceptable risk of error. Therefore, the raw material base will be of key importance to ensure access to wind energy at the level adopted in PEP2040. This is especially true for REE, which has limited global resources and most of its imports come from China. Therefore, REE-related variables turned out to be a destimulant. This means that the growing demand for REE may pose a threat to the further development of wind farms. The destimulant is also an explanatory variable that characterizes manganese consumption. In this case, the growing global demand may hamper the development of wind energy in Poland. It should be remembered that manganese is not mined in Poland, but is imported, which may pose a threat to the development of wind energy. Poland is not the only country that in the coming years will strive to diversify the sources of electricity and thus develop renewable energy in the form of wind energy. It will be a strategy implemented both in other EU countries and in other countries around the world. As a result, access to raw materials may become increasingly difficult, and their prices may increase. Due to the great importance of critical mineral resources for the further development of wind energy and ensuring Poland’s energy security, it is advisable to look for new sources of obtaining them. In the case of most of the critical raw materials necessary for the production of wind technology, it is possible to recover them from mining waste and waste produced during the coal combustion process. This applies to rare earth elements, which can be recovered from fly ash and manganese, which, for example, in the Czech Republic will be recovered from flotation tailings.

The main limitation of the research is the inability to include qualitative data in the proposed solution. The next step of the research will be the use of fuzzy sets, which will allow the analysis to include information on public opinion about the impact of wind farms on the natural environment as well as to take into account the legal factors that greatly influence the pace of development of wind energy.

Symbols

- **ARMAX** – Autoregressive Moving Average with Exogenous Input model,
- **VIF** – Variance Inflation Factor,
- **MAPE** – Mean Absolute Percentage Error,
- **PEP2040** – Polish Energy Policy until 2040,
- **REE** – rare earth elements,
- **α** – significance level,
- **y(t)** – output signal sequence,
- **t** – time,
- **a_i, b_j, c_k** – prediction coefficients,
- **e(t)** – white noise,
- **u(t)** – input signal string,
- **n** – predictor order,
- **i, j, k** – lag,
\( R_j^2 \) – determination coefficient,
\( \delta \delta \) – deterministic trend coefficient,
\( DF \) – Dickey-Fuller test,
\( BIC \) – Schwarz information criterion,
\( HQ \) – Hannan-Quinn information criterion,
\( AIC \) – Akaike information criterion,
\( K \) – number of model parameters,
\( L(\hat{\theta}) \) – the credibility function of the model corrected by the penalty function,
\( e_t \) – prediction error,
\( RMSE \) – Root Mean Square error,
\( MAE \) – Mean Absolute error,
\( \hat{y}_t \) – forecast of the dependent variable in period \( t \),
\( U \) – Theil’s inequality Coefficient,
\( Q \) – Ljung Box test,
\( \hat{\rho}_k^2 \) – sample autocorrelation at lag \( k \),
\( h \) – number of lags tested,
\( DH \) – Doornik-Hansen test,
\( Z(\sqrt{h}) \) – transformed sample skewness,
\( z_2 \) – Wilson-Hilferty transformation,
\( LCOE \) – levelized cost of electricity,
\( p \)-value – probability of obtaining the observed results.

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REFERENCES


The Impact of Selected Factors and Access to Mineral Resources on the Development of Wind Energy in Poland

Keywords
wind energy, ARMAX model, forecasting, rare earth metals

Abstract
This article presents the results of research on the importance of access to critical raw materials for the development of wind energy in Poland. The authors have built a set of factors that can potentially influence this development. Twenty-four explanatory variables were taken into account, which were assigned to five categories. The amount of demand for mineral resources related to the development of wind technology was determined using a computer program written by the authors. The importance of individual factors was verified using the ARMAX model. As a result of this, it was possible to identify the explanatory variables that significantly affect the volume of wind energy production in Poland. The group of mineral resources includes critical metals that are necessary for the production of wind turbines. These are rare earth elements, copper, nickel, boron and manganese. The ARMAX model enables the examination of the relationship between the explained variable and the explanatory variables. Optimization of the model parameters was performed by limiting the mean square error. During the validation of the model, the VIF (variance inflation factor), Dickey-Fuller and Doornik-Hansen tests were used. The ARMAX validation also consisted of selecting the model characterized by the lowest value of information criteria and determining ex post errors, including the mean absolute percentage error (MAPE). In addition, the nature of individual independent variables was determined, i.e. whether they were stimulants, nominants, or destimulants. The forecast made it possible to verify the possibility of meeting the assumptions of the Polish Energy Policy until 2040. It showed that if the development trends of the factors that affect wind energy do not change, it would be possible to meet the assumptions of PEP2040 regarding the dynamic development of wind farms in Poland and double the generation capacity by 2030. Analysis using the ARMAX model showed that access to raw materials such as REE, Cu, Ni, Br and Mn would have a very significant impact on the development of wind energy in Poland. Each factor of the raw material category that was introduced into the model was considered statistically significant at the significance level of $\alpha = 0.01$, i.e. at the lowest acceptable risk of error. Therefore, the raw material base would be of key importance to ensure access to wind energy at the level adopted in PEP2040.
WPŁYW WYBRANYCH CZYNNIKÓW I DOSTĘPU DO SUROWCÓW MINERALNYCH NA ROZWÓJ ENERGETYKI WIAТROWEJ W POLSCE

Słowa kluczowe
energia wiatrowa, model ARMAX, prognozowanie, metale ziem rzadkich

Streszczenie